

# A PLANE IN MDS CONFIGURATION SPACE

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## 1. PROBLEM

The *Multidimensional Scaling* or *MDS* problem (of locating  $n$  point in  $p$  dimensions) is to minimize the function

$$(1) \quad \sigma(X) = \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n w_{ij} (\delta_{ij} - d_{ij}(X))^2$$

over all  $n \times p$  matrices  $X$ . Here  $w_{ij}$  are known *weights* and  $\delta_{ij}$  are known *dissimilarities*. ToAlso

$$d_{ij}(X) = \|\mathbf{x}_i - \mathbf{x}_j\| = \sqrt{(\mathbf{x}_i - \mathbf{x}_j)'(\mathbf{x}_i - \mathbf{x}_j)}$$

is the Euclidean distance between rows  $i$  and  $j$  of  $X$ .

## 2. TRANSFORMATION

By choosing a basis  $Y_1, \dots, Y_m$  for the space of all  $n \times p$  matrices  $\mathbb{R}^{n \times p}$ , we can write  $X$  in the form

$$(2) \quad X = \sum_{v=1}^m \gamma_v Y_v.$$

This also makes it possible to consider the more general problem in which  $X$  varies in an  $m$ -dimensional subspace of  $\mathbb{R}^{n \times p}$ . The MDS problem now becomes minimization of the loss function  $\sigma(\gamma)$  over the coefficients in (2).

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Let us introduce some notation to simplify this problem [De Leeuw, 1993]. We can write

$$d_{ij}^2(\boldsymbol{y}) = \mathbf{tr}\left(\sum_{\nu=1}^m \boldsymbol{y}_\nu Y_\nu\right)' A_{ij} \left(\sum_{\nu=1}^m \boldsymbol{y}_\nu Y_\nu\right),$$

where

$$A_{ij} = (\boldsymbol{e}_i - \boldsymbol{e}_j)(\boldsymbol{e}_i - \boldsymbol{e}_j)'$$

and  $\boldsymbol{e}_i$  and  $\boldsymbol{e}_j$  are unit vectors (columns of the identity matrix). Define the  $m \times m$  matrices  $C_{ij}$ , with elements

$$(C_{ij})_{\nu\mu} = \mathbf{tr} Y_\nu' A_{ij} Y_\mu.$$

Observe the  $C_{ij}$  are positive semi-definite. Now

$$d_{ij}^2(\boldsymbol{y}) = \boldsymbol{y}' C_{ij} \boldsymbol{y}$$

Suppose, without loss of generality, that the basis  $Y_\nu$  is chosen in such a way that

$$\sum_{i=1}^n \sum_{j=1}^n w_{ij} C_{ij} = I.$$

Then

$$(3) \quad \sigma(\boldsymbol{y}) = 1 + \frac{1}{2} \boldsymbol{y}' \boldsymbol{y} - \sum_{i=1}^n \sum_{j=1}^n w_{ij} \delta_{ij} \sqrt{\boldsymbol{y}' C_{ij} \boldsymbol{y}},$$

where we have also assumed, without loss of generality, that

$$\frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n w_{ij} \delta_{ij}^2 = 1.$$

The MDS problem, which we call problem  $\mathcal{P}$  from now on, is clearly equivalent to minimizing (3) over  $\boldsymbol{y}$ .

### 3. NECESSARY CONDITIONS FOR A LOCAL MINIMUM

In a neighborhood of each local minimum we have  $d_{ij}(\boldsymbol{y}) > 0$  for all  $i \neq j$  [De Leeuw, 1984]. Thus the loss function is differentiable at local minima and

$$\frac{\partial \sigma}{\partial \boldsymbol{y}} = (I - B(\boldsymbol{y})) \boldsymbol{y},$$

with

$$B(\boldsymbol{y}) = \sum_{i=1}^n \sum_{j=1}^n w_{ij} \frac{\delta_{ij}}{d_{ij}(\boldsymbol{y})} C_{ij}.$$

Also

$$\frac{\partial^2 \sigma}{\partial \boldsymbol{y} \partial \boldsymbol{y}} = I - H(\boldsymbol{y}),$$

where

$$H(\boldsymbol{y}) = \sum_{i=1}^n \sum_{j=1}^n w_{ij} \frac{\delta_{ij}}{d_{ij}(\boldsymbol{y})} \left[ C_{ij} - \frac{C_{ij} \boldsymbol{y} \boldsymbol{y}' C_{ij}}{\boldsymbol{y}' C_{ij} \boldsymbol{y}} \right].$$

Thus at local minimum we must have

$$(4a) \quad B(\boldsymbol{y}) \boldsymbol{y} = \boldsymbol{y},$$

and in addition

$$(4b) \quad H(\boldsymbol{y}) \preceq I.$$

Or, in words,  $\boldsymbol{y}$  must be an eigenvector of  $B(\boldsymbol{y})$  corresponding with an eigenvalue equal to one, and all eigenvalues of  $H(\boldsymbol{y})$  must be less than or equal to one. Observe that both  $B(\boldsymbol{y})$  and  $H(\boldsymbol{y})$  are positive semi-definite. Because  $H(\boldsymbol{y}) \boldsymbol{y} = 0$  the matrix  $H(\boldsymbol{y})$  is always singular. Also note that if  $H(\boldsymbol{y}) < I$  then  $\boldsymbol{y}$  is a local minimum.

The majorization algorithm [De Leeuw, 1977] for MDS takes the simple form

$$\boldsymbol{y}^{(k+1)} = B(\boldsymbol{y}^{(k)}) \boldsymbol{y}^{(k)},$$

while the Newton-Raphson method can be written in the form

$$\boldsymbol{y}^{(k+1)} = (I - H(\boldsymbol{y}^{(k)}))^{-1} B(\boldsymbol{y}^{(k)}) \boldsymbol{y}^{(k)}.$$

#### 4. RELAXATION

There is yet another way of rewriting the problem. Define

$$(5) \quad \sigma(\Gamma) = 1 + \frac{1}{2} \mathbf{tr} \Gamma - \sum_{i=1}^n \sum_{j=1}^n w_{ij} \delta_{ij} \sqrt{\mathbf{tr} C_{ij} \Gamma},$$

Problem  $\mathcal{P}$  is now minimizing (5) over all positive semi-definite  $\Gamma$  of rank one, i.e. all  $\Gamma$  of the form  $\Gamma = \boldsymbol{y} \boldsymbol{y}'$ .

Now define problem  $\bar{\mathcal{P}}$ , which is the problem of minimizing (5) over all positive semi-definite matrices  $\Gamma$ , without imposing the rank one constraint. Problem  $\bar{\mathcal{P}}$  is a *convex relaxation* of problem  $\mathcal{P}$ , because the loss function (5) is convex in  $\Gamma$  and in  $\bar{\mathcal{P}}$  we require  $\Gamma$  to vary in a convex set. The necessary and sufficient conditions for a local (and thus global) minimum in  $\bar{\mathcal{P}}$  are

$$(6a) \quad I - B(\Gamma) \succeq 0,$$

$$(6b) \quad \Gamma \succeq 0,$$

$$(6c) \quad \mathbf{tr} \Gamma(I - B(\Gamma)) = 0,$$

where

$$B(\Gamma) = \sum_{i=1}^n \sum_{j=1}^n w_{ij} \frac{\delta_{ij}}{d_{ij}(\Gamma)} C_{ij},$$

and  $d_{ij}(\Gamma) = \sqrt{\mathbf{tr} C_{ij}\Gamma}$ .

If the solution  $\hat{\Gamma}$  of  $\bar{\mathcal{P}}$  is of rank one, then  $\hat{\Gamma} = \hat{y}\hat{y}'$ , and obviously  $\hat{y}$  solves  $\mathcal{P}$ . This means that not only (4a) and (4b) are satisfied, but actually  $\hat{y}$  gives the global minimum of the loss function (3).

If  $\hat{y}$  is a solution of (4a), then  $\hat{\Gamma} = \hat{y}\hat{y}'$  satisfies both (6b) and (6c). If  $B(\hat{y}) \preceq I$ , i.e. if  $\hat{y}$  is the eigenvector of  $B(\hat{y})$  corresponding to the largest eigenvalue, then (6a) is satisfied as well. Thus  $\hat{\Gamma}$  solves  $\bar{\mathcal{P}}$ , and  $\hat{y}$  gives the global minimum of (3). Observe that  $H(y) \preceq B(y)$  for all  $y$ , which means that for loss function (3) the sufficient condition for a local minimum  $H(y) < I$  is weaker than the sufficient condition for global minimum  $B(y) \preceq I$  (as it should be, of course).

## 5. A PLANE IN CONFIGURATION SPACE

We can obtain some additional results if  $m = 2$ , i.e. if  $X$  is in a plane in  $\mathbb{R}^{n \times p}$  spanned by just two matrices  $Y_1$  and  $Y_2$ . In that case the solution  $\hat{\Gamma}$  of  $\bar{\mathcal{P}}$  is a matrix of order two. The rank of  $\hat{\Gamma}$  can be either zero, one, or two. We can safely exclude  $\Gamma = 0$ , because we know that at a local minimum all  $d_{ij}(\Gamma)$  must be positive. If

the rank of  $\hat{\Gamma}$  is two, then  $B(\hat{\Gamma}) = I$ . Because the  $C_{ij}$  are linearly independent this implies that  $d_{ij}(\hat{\Gamma}) = \delta_{ij}$ , i.e.  $\sigma(\hat{\Gamma}) = 0$ . It follows that if  $\sigma(\hat{\Gamma}) > 0$  then  $\hat{\Gamma}$  is of rank one and gives the global minimum  $\hat{y}$  of  $\mathcal{P}$ .

$B(\gamma)\gamma = \gamma$  means that  $\gamma$  is an eigenvector of both  $B(\gamma)$  and  $H(\gamma)$ , with eigenvalues one and zero, respectively. If  $m = 2$  then  $\bar{y}$ , which is orthogonal to  $\gamma$ , is also an eigenvector of both  $B(\gamma)$  and  $H(\gamma)$ , with eigenvalues, say,  $\mu$  and  $\xi$ . We know that  $0 \leq \xi \leq \mu$ . If  $0 \leq \xi \leq 1 \leq \mu$  we have a local minimum, if  $0 \leq \xi \leq \mu \leq 1$  we have a global minimum.

#### REFERENCES

- J. De Leeuw. Differentiability of Kruskal's Stress at a Local Minimum. *Psychometrika*, 49:111-113, 1984.
- J. De Leeuw. Fitting Distances by Least Squares. Technical Report UCLA Statistics Series 130, Interdivisional Program in Statistics, UCLA, Los Angeles, California, 1993.
- J. De Leeuw. Applications of Convex Analysis to Multidimensional Scaling. In J.R. Barra, F. Brodeau, G. Romier, and B. Van Cutsem, editors, *Recent developments in statistics*, pages 133-145, Amsterdam, The Netherlands, 1977. North Holland Publishing Company.

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